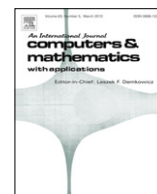


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## Computers and Mathematics with Applications

journal homepage: [www.elsevier.com/locate/camwa](http://www.elsevier.com/locate/camwa)Two-stage least squares based iterative identification algorithm for controlled autoregressive moving average (CARMA) systems<sup>☆</sup>Guoyu Yao<sup>a</sup>, Ruifeng Ding<sup>b,\*</sup><sup>a</sup> Key Laboratory of Advanced Process Control for Light Industry (Ministry of Education), Jiangnan University, Wuxi 214122, PR China<sup>b</sup> School of Internet of Things Engineering, Jiangnan University, Wuxi 214122, PR China

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## ABSTRACT

A two-stage least squares based iterative (two-stage LSI) identification algorithm is derived for controlled autoregressive moving average (CARMA) systems. The basic idea is to decompose a CARMA system into two subsystems and to identify each subsystem, respectively. Because the dimensions of the involved covariance matrices in each subsystem become small, the proposed algorithm has a high computational efficiency. The simulation results indicate that the proposed algorithm is effective.

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## 1. Introduction

The iterative numerical algorithms can be used to solve not only some matrix equations [1–9], but also parameter estimation and filtering problems [10–18]. The iterative identification algorithms make sufficient use of all input–output data and can improve the parameter estimation accuracy [19–21]. For example, Ding et al. presented least squares based iterative algorithms for Hammerstein nonlinear ARMAX systems [22], and for OE and OEMA systems [19]; Han et al. gave a hierarchical least squares based iterative identification algorithm for a class of multivariable CARMA-like systems [23]. Zhang et al. proposed a hierarchical gradient based iterative parameter estimation algorithm for multivariable output error moving average systems (i.e., multivariable OEMA-like systems) [24]. Bao et al. developed a least squares based iterative identification method for multivariable controlled ARMA systems [25]. Ding et al. presented a least squares based iterative algorithm for controlled autoregressive autoregressive moving average (CARARMA) systems [26]. Also, a least squares based iterative algorithm and a gradient based iterative algorithm are developed for Box–Jenkins systems [20,21].

Two-stage algorithms have been widely used in the identification field. Bai presented an optimal two-stage identification algorithm for Hammerstein–Wiener nonlinear systems in the sense of a weighted nonlinear least squares criterion [27]. Li et al. gave a two-stage algorithm identification of nonlinear dynamic systems to reduce computational complexity [28]. Hwang et al. developed a two-stage least squares algorithm to identify continuous systems with time delay based on the pulse responses [29]. Cao gave a rounding error analysis of two-stage iterative methods for large linear systems [30].

The least squares based iterative algorithm was used to identify the Box–Jenkins models with finite measurement data [21], but it requires large computational load, especially when the dimension of the involved covariance matrices is large. Duan et al. presented a two-stage recursive least squares parameter estimation algorithm for output error models [31]. This paper uses the decomposition technique and proposes a two-stage iterative identification method for controlled

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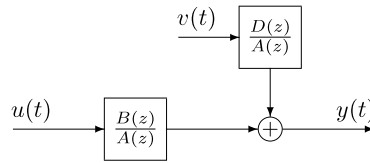


Fig. 1. A system described by the CARMA model.

autoregressive moving average (CARMA) systems. The basic idea is to decompose a CARMA system into two subsystems, to transform the original identification problem into two subproblems with small dimensions, and to identify the parameters of each subsystem, respectively. The proposed algorithm has a high computational efficiency and can be extended to nonlinear systems [32,33].

For several decades, many recursive and iterative parameter estimation algorithms have been developed, including the least squares-based iterative estimation for output error moving average (OEMA) systems using data filtering [34], the input–output data filtering based recursive least squares parameter estimation for CARARMA systems [35], the auxiliary model based least squares or gradient estimation algorithms [36–42], the multi-innovation parameter estimation algorithm for linear and pseudo-linear regression systems [43–52], the hierarchical least squares or hierarchical gradient parameter estimation algorithms [53–63], and other parameter estimation algorithms for linear regressive models [64–73], the recursive and iterative identification methods for Hammerstein systems [74], the identification algorithms for Hammerstein OEMA systems, Hammerstein OEAR systems and Wiener systems [75–80].

The rest of the paper is organized as follows. Section 2 derives a two-stage least squares based iterative identification algorithm for CARMA systems. Section 3 gives the least squares based iterative algorithm for comparisons. Section 4 provides a simulation example to illustrate the effectiveness of the proposed algorithms. Finally, we offer some concluding remarks in Section 5.

## 2. The two-stage least squares algorithm

Let us introduce some notations first. “ $\mathbf{A} \equiv: \mathbf{X}$ ” or “ $\mathbf{X} \equiv: \mathbf{A}$ ” stands for “ $\mathbf{A}$  is defined as  $\mathbf{X}$ ”; the symbol  $\mathbf{I}_n$  stands for an identity matrix of appropriate size ( $n \times n$ );  $\mathbf{1}_n$  represents an  $n$ -dimensional column vector whose elements are all 1; the superscript T denotes the matrix transpose; the norm of a matrix  $\mathbf{X}$  is defined by  $\|\mathbf{X}\|^2 = \text{tr}[\mathbf{X}\mathbf{X}^T]$ .

Consider the CARMA system, depicted in Fig. 1,

$$A(z)y(t) = B(z)u(t) + D(z)v(t), \quad (1)$$

where  $\{u(t)\}$  and  $\{y(t)\}$  are the input and output sequences of the system, respectively,  $\{v(t)\}$  is a white noise sequence with zero mean and variance  $\sigma^2$ , and  $A(z)$ ,  $B(z)$  and  $D(z)$  are the polynomials, of known orders ( $n_a$ ,  $n_b$ ,  $n_d$ ), in the unit backward shift operator  $z^{-1}$  [i.e.,  $z^{-1}y(t) = y(t-1)$ ], and defined by

$$A(z) := 1 + a_1z^{-1} + a_2z^{-2} + \cdots + a_{n_a}z^{-n_a},$$

$$B(z) := b_1z^{-1} + b_2z^{-2} + \cdots + b_{n_b}z^{-n_b},$$

$$D(z) := 1 + d_1z^{-1} + d_2z^{-2} + \cdots + d_{n_d}z^{-n_d}.$$

Without loss of generality, assume that  $y(t) = 0$ ,  $u(t) = 0$  and  $v(t) = 0$  for  $t \leq 0$ .

The objective of this paper is to apply the decomposition technique and derive a two-stage least squares based iterative identification algorithm for estimating the system parameters  $a_i$ ,  $b_i$  and  $d_i$ .

Define the system parameter vectors,

$$\Theta := \begin{bmatrix} \theta \\ \vartheta \end{bmatrix} \in \mathbb{R}^n, \quad n := n_a + n_b + n_d,$$

$$\theta := [a_1, a_2, \dots, a_{n_a}, b_1, b_2, \dots, b_{n_b}]^T \in \mathbb{R}^{n_a+n_b},$$

$$\vartheta := [d_1, d_2, \dots, d_{n_d}]^T \in \mathbb{R}^{n_d},$$

and the information vectors,

$$\varphi(t) := \begin{bmatrix} \phi(t) \\ \psi(t) \end{bmatrix} \in \mathbb{R}^n,$$

$$\phi(t) := [-y(t-1), -y(t-2), \dots, -y(t-n_a), u(t-1), u(t-2), \dots, u(t-n_b)]^T \in \mathbb{R}^{n_a+n_b}, \quad (2)$$

$$\psi(t) := [v(t-1), v(t-2), \dots, v(t-n_d)]^T \in \mathbb{R}^{n_d}. \quad (3)$$

From (1), we can obtain the following identification model:

$$y(t) = \phi^T(t)\theta + \psi^T(t)\vartheta + v(t) \quad (4)$$

$$= \varphi^T(t)\Theta + v(t). \quad (5)$$

Define two intermediate variables:

$$y_1(t) := y(t) - \boldsymbol{\psi}^T(t)\boldsymbol{\vartheta}, \quad (6)$$

$$y_2(t) := y(t) - \boldsymbol{\phi}^T(t)\boldsymbol{\theta}. \quad (7)$$

From (6) and (7), the system in (4) can be decomposed into two “fictitious” subsystems:

$$y_1(t) = \boldsymbol{\phi}^T(t)\boldsymbol{\theta} + v(t), \quad (8)$$

$$y_2(t) = \boldsymbol{\psi}^T(t)\boldsymbol{\vartheta} + v(t). \quad (9)$$

Consider the data from  $t = 1$  to  $t = L$  ( $L \gg n$ ) and define the stacked output vectors  $\mathbf{Y}(L)$ ,  $\mathbf{Y}_1(L)$  and  $\mathbf{Y}_2(L)$ , the stacked information matrices  $\boldsymbol{\Phi}(L)$  and  $\boldsymbol{\Psi}(L)$  and the stacked white noise vector  $\mathbf{V}(L)$  as

$$\begin{aligned} \mathbf{Y}(L) &:= \begin{bmatrix} y(1) \\ y(2) \\ \vdots \\ y(L) \end{bmatrix} \in \mathbb{R}^L, & \mathbf{V}(L) &:= \begin{bmatrix} v(1) \\ v(2) \\ \vdots \\ v(L) \end{bmatrix} \in \mathbb{R}^L, \\ \mathbf{Y}_1(L) &:= \begin{bmatrix} y_1(1) \\ y_1(2) \\ \vdots \\ y_1(L) \end{bmatrix} \in \mathbb{R}^L, & \mathbf{Y}_2(L) &:= \begin{bmatrix} y_2(1) \\ y_2(2) \\ \vdots \\ y_2(L) \end{bmatrix} \in \mathbb{R}^L, \\ \boldsymbol{\Phi}(L) &:= \begin{bmatrix} \boldsymbol{\phi}^T(1) \\ \boldsymbol{\phi}^T(2) \\ \vdots \\ \boldsymbol{\phi}^T(L) \end{bmatrix} \in \mathbb{R}^{L \times (n_a + n_b)}, & \boldsymbol{\Psi}(L) &:= \begin{bmatrix} \boldsymbol{\psi}^T(1) \\ \boldsymbol{\psi}^T(2) \\ \vdots \\ \boldsymbol{\psi}^T(L) \end{bmatrix} \in \mathbb{R}^{L \times n_d}. \end{aligned}$$

Note that  $\mathbf{Y}(L)$ ,  $\boldsymbol{\Phi}(L)$  and  $\boldsymbol{\Psi}(L)$  contain all the measured data  $\{u(t), y(t) : t = 1, 2, \dots, L\}$ . From (6) and (7), we have

$$\mathbf{Y}_1(L) = \mathbf{Y}(L) - \boldsymbol{\Psi}(L)\boldsymbol{\vartheta}, \quad (10)$$

$$\mathbf{Y}_2(L) = \mathbf{Y}(L) - \boldsymbol{\Phi}(L)\boldsymbol{\theta}. \quad (11)$$

From (8) and (9), we have

$$\mathbf{Y}_1(L) = \boldsymbol{\Phi}(L)\boldsymbol{\theta} + \mathbf{V}(L),$$

$$\mathbf{Y}_2(L) = \boldsymbol{\Psi}(L)\boldsymbol{\vartheta} + \mathbf{V}(L).$$

Define two quadratic criterion functions:

$$J_1(\boldsymbol{\theta}) := \|\mathbf{Y}_1(L) - \boldsymbol{\Phi}(L)\boldsymbol{\theta}\|^2,$$

$$J_2(\boldsymbol{\vartheta}) := \|\mathbf{Y}_2(L) - \boldsymbol{\Psi}(L)\boldsymbol{\vartheta}\|^2.$$

For these two optimization problems, letting the partial derivatives of  $J_1(\boldsymbol{\theta})$  and  $J_2(\boldsymbol{\vartheta})$  with respect to  $\boldsymbol{\theta}$  and  $\boldsymbol{\vartheta}$  be zero gives

$$\frac{\partial J_1(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = -2\boldsymbol{\Phi}^T(L)[\mathbf{Y}_1(L) - \boldsymbol{\Phi}(L)\boldsymbol{\theta}] = \mathbf{0},$$

$$\frac{\partial J_2(\boldsymbol{\vartheta})}{\partial \boldsymbol{\vartheta}} = -2\boldsymbol{\Psi}^T(L)[\mathbf{Y}_2(L) - \boldsymbol{\Psi}(L)\boldsymbol{\vartheta}] = \mathbf{0}.$$

Assume that the information vectors  $\boldsymbol{\phi}(t)$  and  $\boldsymbol{\psi}(t)$  are persistently exciting, that is,  $[\boldsymbol{\Phi}^T(L)\boldsymbol{\Phi}(L)]$  and  $[\boldsymbol{\Psi}^T(L)\boldsymbol{\Psi}(L)]$  are non-singular. From the above two equations, we can obtain the following least squares estimates of the parameter vectors  $\boldsymbol{\theta}$  and  $\boldsymbol{\vartheta}$ :

$$\hat{\boldsymbol{\theta}} = [\boldsymbol{\Phi}^T(L)\boldsymbol{\Phi}(L)]^{-1}\boldsymbol{\Phi}^T(L)\mathbf{Y}_1(L), \quad (12)$$

$$\hat{\boldsymbol{\vartheta}} = [\boldsymbol{\Psi}^T(L)\boldsymbol{\Psi}(L)]^{-1}\boldsymbol{\Psi}^T(L)\mathbf{Y}_2(L). \quad (13)$$

Substituting (10) into (12) and (11) into (13) gives

$$\hat{\boldsymbol{\theta}} = [\boldsymbol{\Phi}^T(L)\boldsymbol{\Phi}(L)]^{-1}\boldsymbol{\Phi}^T(L)[\mathbf{Y}(L) - \boldsymbol{\Psi}(L)\boldsymbol{\vartheta}], \quad (14)$$

$$\hat{\boldsymbol{\vartheta}} = [\boldsymbol{\Psi}^T(L)\boldsymbol{\Psi}(L)]^{-1}\boldsymbol{\Psi}^T(L)[\mathbf{Y}(L) - \boldsymbol{\Phi}(L)\boldsymbol{\theta}]. \quad (15)$$

However, the right-hand sides of (14) and (15) contain the unknown parameter  $\boldsymbol{\vartheta}$  and  $\boldsymbol{\theta}$ , respectively, it is impossible to compute the estimates  $\hat{\boldsymbol{\theta}}$  and  $\hat{\boldsymbol{\vartheta}}$ . The approach is based on the hierarchical identification principle: let  $k = 1, 2, 3, \dots$  be an

iteration variable,  $\hat{\Theta}_k := \begin{bmatrix} \hat{\theta}_k \\ \hat{\vartheta}_k \end{bmatrix}$  be the iterative estimate of  $\Theta = \begin{bmatrix} \theta \\ \vartheta \end{bmatrix}$  at iteration  $k$ , and  $\hat{v}_k(t)$  be the estimate of  $v(t)$  at iteration  $k$ , and define

$$\begin{aligned} \hat{\phi}_k(t) &:= \begin{bmatrix} \phi(t) \\ \hat{\psi}_k(t) \end{bmatrix} \in \mathbb{R}^{n_a+n_b+n_d}, \\ \hat{\psi}_k(t) &:= [\hat{v}_{k-1}(t-1), \hat{v}_{k-1}(t-2), \dots, \hat{v}_{k-1}(t-n_d)]^T \in \mathbb{R}^{n_d}, \\ \hat{\Psi}_k(L) &:= \begin{bmatrix} \hat{\psi}_k^T(1) \\ \hat{\psi}_k^T(2) \\ \vdots \\ \hat{\psi}_k^T(L) \end{bmatrix} \in \mathbb{R}^{L \times n_d}. \end{aligned}$$

From (4), we have

$$v(t) = y(t) - \phi^T(t)\theta - \psi^T(t)\vartheta.$$

Replacing  $\psi(t)$ ,  $\theta$  and  $\vartheta$  with  $\hat{\psi}_k(t)$ ,  $\hat{\theta}_k$  and  $\hat{\vartheta}_k$ , the estimate  $\hat{v}_k(t)$  of  $v(t)$  can be computed by

$$\begin{aligned} \hat{v}_k(t) &= y(t) - \phi^T(t)\hat{\theta}_k - \hat{\psi}_k^T(t)\hat{\vartheta}_k \\ &= y(t) - \hat{\phi}_k^T(t)\hat{\Theta}_k. \end{aligned} \quad (16)$$

Replacing  $\Psi(L)$  and  $\vartheta$  in (14) with their estimates  $\hat{\Psi}_k(L)$  and  $\hat{\vartheta}_{k-1}$ , and replacing  $\Psi(L)$  and  $\theta$  in (15) with their estimates  $\hat{\Psi}_k(L)$  and  $\hat{\theta}_k$ , we can summarize the two-stage least squares based iterative (two-stage LSI) identification algorithm for estimating  $\theta$  and  $\vartheta$  of the CARMA systems as follows:

$$\hat{\theta}_k = [\Phi^T(L)\Phi(L)]^{-1}\Phi^T(L)[Y(L) - \hat{\Psi}_k(L)\hat{\vartheta}_{k-1}], \quad (17)$$

$$\hat{\vartheta}_k = [\hat{\Psi}_k^T(L)\hat{\Psi}_k(L)]^{-1}\hat{\Psi}_k^T(L)[Y(L) - \Phi(L)\hat{\theta}_k], \quad k = 1, 2, \dots, \quad (18)$$

$$Y(L) = [y(1), y(2), \dots, y(L)]^T, \quad (19)$$

$$\Phi(L) = [\phi(1), \phi(2), \dots, \phi(L)]^T, \quad (20)$$

$$\hat{\Psi}_k(L) = [\hat{\psi}_k(1), \hat{\psi}_k(2), \dots, \hat{\psi}_k(L)]^T, \quad (21)$$

$$\phi(t) = [-y(t-1), -y(t-2), \dots, -y(t-n_a), u(t-1), u(t-2), \dots, u(t-n_b)]^T, \quad (22)$$

$$\hat{\psi}_k(t) = [\hat{v}_{k-1}(t-1), \hat{v}_{k-1}(t-2), \dots, \hat{v}_{k-1}(t-n_d)]^T, \quad (23)$$

$$\hat{\phi}_k(t) = [\phi^T(t), \hat{\psi}_k^T(t)]^T, \quad (24)$$

$$\hat{\Theta}_k := [\hat{\theta}_k^T, \hat{\vartheta}_k^T]^T, \quad (25)$$

$$\hat{v}_k(t) = y(t) - \hat{\phi}_k^T(t)\hat{\Theta}_k. \quad (26)$$

In the above algorithm, the dimension of the covariance matrix  $S_1 := [\Phi^T(L)\Phi(L)]^{-1}$  in (17) is  $(n_a + n_b) \times (n_a + n_b)$ ,  $S_2 := [\hat{\Psi}_k^T(L)\hat{\Psi}_k(L)]^{-1}$  in (18) is  $n_d \times n_d$ .

The steps involved in the two-stage LSI algorithm in (17)–(26) to compute the parameter estimates  $\hat{\theta}_k$  and  $\hat{\vartheta}_k$  for CARMA systems are listed below.

1. Collect the input–output data  $\{u(t), y(t) : i = 1, 2, \dots, L\}$  ( $L$  is the data length), form  $Y(L)$  by (19),  $\phi(t)$  by (22) and  $\Phi(L)$  by (20), and give the parameter estimation precision  $\varepsilon = 0.01$ .
2. To initialize, let  $k = 1$ ,  $\hat{\vartheta}_0 = \mathbf{1}_{n_d}/p_0$ ,  $p_0 = 10^6$ ,  $\hat{v}_0(t)$  = a random number.
3. Form  $\hat{\psi}_k(t)$  by (23),  $\hat{\Psi}_k(L)$  by (21) and  $\hat{\phi}_k(t)$  by (24).
4. Update  $\hat{\theta}_k$  by (17) and  $\hat{\vartheta}_k$  by (18).
5. Form  $\hat{\Theta}_k$  by (25).
6. Compute  $\hat{v}_k(t)$  by (26).
7. If  $\|\hat{\Theta}_k - \Theta\| = \|\hat{\theta}_k - \hat{\theta}_{k-1}\| + \|\hat{\vartheta}_k - \hat{\vartheta}_{k-1}\| \leq \varepsilon$ , obtain the parameter estimates  $\hat{\theta}_k$  and  $\hat{\vartheta}_k$ ; otherwise, increase  $k$  by 1 and go to step 3.

The flowchart of computing the parameter estimates  $\hat{\theta}_k$  and  $\hat{\vartheta}_k$  is shown in Fig. 2.

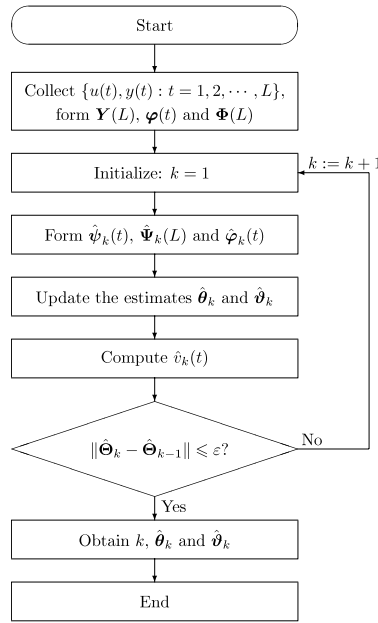


Fig. 2. The flowchart for computing the two-stage LSI parameter estimates.

### 3. The least squares based iterative algorithm

To show the advantages of the proposed two-stage LSI algorithm, the following gives the least squares based iterative (LSI) algorithm for comparisons. Define the stacked information matrix

$$\mathbf{H}(L) := \begin{bmatrix} \boldsymbol{\varphi}^T(1) \\ \boldsymbol{\varphi}^T(2) \\ \vdots \\ \boldsymbol{\varphi}^T(L) \end{bmatrix} \in \mathbb{R}^{L \times (n_a + n_b + n_d)}.$$

Using the definitions of  $\mathbf{Y}(L)$ ,  $\mathbf{V}(L)$  and  $\mathbf{H}(L)$ , Eq. (5) can be written as

$$\mathbf{Y}(L) = \mathbf{H}(L)\boldsymbol{\Theta} + \mathbf{V}(L).$$

Define a quadratic criterion function:

$$J_3(\boldsymbol{\Theta}) := \|\mathbf{Y}(L) - \mathbf{H}(L)\boldsymbol{\Theta}\|^2.$$

Minimizing the criterion function  $J_3(\boldsymbol{\Theta})$  and letting its partial derivative with respect to  $\boldsymbol{\Theta}$  be zero, we can obtain the least squares algorithm for estimating the parameter vector  $\boldsymbol{\Theta}$ :

$$\hat{\boldsymbol{\Theta}} = [\mathbf{H}^T(L)\mathbf{H}(L)]^{-1}\mathbf{H}^T(L)\mathbf{Y}(L). \quad (27)$$

Here, a difficulty arises because  $\mathbf{H}(L)$  in the above equation (that is  $\boldsymbol{\varphi}(t)$  in (3)) contains the unknown noise terms  $v(t-i)$ . The approach is based on the hierarchical identification principle: let  $\hat{\boldsymbol{\varphi}}_k(t)$  denote the information vector by replacing  $v(t-i)$  in (3) with their estimates  $\hat{v}_{k-1}(t-i)$  at iteration  $k-1$ ,  $\hat{\mathbf{H}}_k(L)$  denote the stacked information matrix obtained by replacing  $\boldsymbol{\varphi}(t-i)$  in  $\mathbf{H}(L)$  with  $\hat{\boldsymbol{\varphi}}_k(t-i)$ , and the estimates  $\hat{v}_k(t-i)$  of  $v(t-i)$  at iteration  $k$  can be computed by (26).

Replacing  $\mathbf{H}(L)$  in (27) with  $\hat{\mathbf{H}}_k(L)$ , we can obtain the least squares based iterative identification algorithm for estimating  $\boldsymbol{\Theta}$  of the CARMA systems as follows [21]:

$$\hat{\boldsymbol{\Theta}}_k = [\hat{\mathbf{H}}_k^T(L)\hat{\mathbf{H}}_k(L)]^{-1}\hat{\mathbf{H}}_k^T(L)\mathbf{Y}(L), \quad (28)$$

$$\mathbf{Y}(L) = [y(1), y(2), \dots, y(L)]^T, \quad (29)$$

$$\hat{\mathbf{H}}_k(L) = [\hat{\boldsymbol{\varphi}}_k(1), \hat{\boldsymbol{\varphi}}_k(2), \dots, \hat{\boldsymbol{\varphi}}_k(L)]^T, \quad (30)$$

$$\begin{aligned} \hat{\boldsymbol{\varphi}}_k(t) = & [-y(t-1), -y(t-2), \dots, -y(t-n_a), u(t-1), u(t-2), \dots, u(t-n_b), \\ & \hat{v}_{k-1}(t-1), \hat{v}_{k-1}(t-2), \dots, \hat{v}_{k-1}(t-n_d)]^T, \end{aligned} \quad (31)$$

$$\hat{v}_k(t) = y(t) - \hat{\boldsymbol{\varphi}}_k^T(t)\hat{\boldsymbol{\Theta}}_k. \quad (32)$$

Compared the LSI algorithm and the two-stage LSI algorithm, we can draw the following conclusions.

**Table 1**The parameter estimates and their errors with  $\sigma^2 = 0.10^2$ .

Algorithms	$t = L$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
LSI	1000	-1.59746	0.79877	0.39895	0.30981	-0.59785	2.20821
	2000	-1.60034	0.80072	0.39985	0.30210	-0.61364	1.34686
	3000	-1.60041	0.80051	0.40074	0.29990	-0.62064	0.98659
Two-stage LSI	1000	-1.60215	0.80342	0.39885	0.30721	-0.59985	2.08739
	2000	-1.60263	0.80298	0.39983	0.30109	-0.61427	1.32626
	3000	-1.60170	0.80203	0.40072	0.29983	-0.62067	0.99376
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

**Table 2**The parameter estimates and their errors with  $\sigma^2 = 0.50^2$ .

Algorithms	$t = L$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
LSI	1000	-1.58859	0.79212	0.39474	0.34862	-0.58797	3.70276
	2000	-1.59951	0.79752	0.39930	0.31149	-0.61263	1.51670
	3000	-1.60004	0.79577	0.40384	0.30051	-0.62025	1.04672
Two-stage LSI	1000	-1.60651	0.80940	0.39447	0.33795	-0.60288	2.77809
	2000	-1.60739	0.80748	0.39914	0.30773	-0.61841	1.28484
	3000	-1.60382	0.80427	0.40353	0.30097	-0.62248	0.95681
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

**Table 3**The parameter estimates and their errors with  $\sigma^2 = 1.00^2$ .

Algorithms	$t = L$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
LSI	1000	-1.58136	0.78563	0.38961	0.39558	-0.57972	5.89917
	2000	-1.59672	0.79103	0.39868	0.32381	-0.60977	2.01938
	3000	-1.59830	0.78690	0.40799	0.30200	-0.61845	1.35311
Two-stage LSI	1000	-1.60536	0.80536	0.38953	0.37929	-0.60125	4.54018
	2000	-1.60376	0.80163	0.39839	0.31993	-0.61456	1.66025
	3000	-1.59906	0.79814	0.40711	0.30542	-0.61766	1.22932
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

In the LSI algorithm, the dimension of the covariance matrix  $S_3 := [\hat{\mathbf{H}}_k^T(L)\hat{\mathbf{H}}_k(L)]^{-1}$  is  $(n_a + n_b + n_d) \times (n_a + n_b + n_d)$  which is larger than those of  $S_1$  and  $S_2$  and thus the two-stage LSI algorithm has a high computational efficiency.

#### 4. Example

Consider the following simulation system:

$$\begin{aligned}
 A(z)y(t) &= B(z)u(t) + D(z)v(t), \\
 A(z) &= 1 + a_1z^{-1} + a_2z^{-2} = 1 - 1.60z^{-1} + 0.80z^{-2}, \\
 B(z) &= b_1z^{-1} + b_2z^{-2} = 0.40z^{-1} + 0.30z^{-2}, \\
 D(z) &= 1 + d_1z^{-1} = 1 - 0.64z^{-1}, \\
 \boldsymbol{\theta} &= [a_1, a_2, b_1, b_2]^T = [-0.60, 0.80, 0.40, 0.30]^T, \\
 \boldsymbol{\vartheta} &= d_1 = -0.64, \\
 \boldsymbol{\Theta} &= [\boldsymbol{\theta}^T, \boldsymbol{\vartheta}^T]^T.
 \end{aligned}$$

In simulation,  $\{u(t)\}$  is taken as an uncorrelated persistent excitation signal sequence with zero mean and unit variance,  $\{v(t)\}$  as a white noise sequence with zero mean and variance  $\sigma^2$ . Applying the LSI and the two-stage LSI algorithm to estimate the parameters of this system, the parameter estimates with different data length and noise variances are shown in Tables 1–3. The estimation errors for the two-stage LSI algorithm versus  $k$  are shown in Tables 4–6 and Figs. 3–5 where  $\delta := \|\hat{\boldsymbol{\Theta}}_k - \boldsymbol{\Theta}\|/\|\boldsymbol{\Theta}\|$ . When  $\sigma^2 = 0.10^2$ ,  $\sigma^2 = 0.50^2$  and  $\sigma^2 = 1.00^2$ , the corresponding noise-to-signal ratios are  $\delta_{ns} = 7.66\%$ ,  $\delta_{ns} = 38.2\%$ , and  $\delta_{ns} = 76.5\%$ , respectively.

From the simulation results in Tables 1–6 and Figs. 3–5, we can draw the following conclusions.

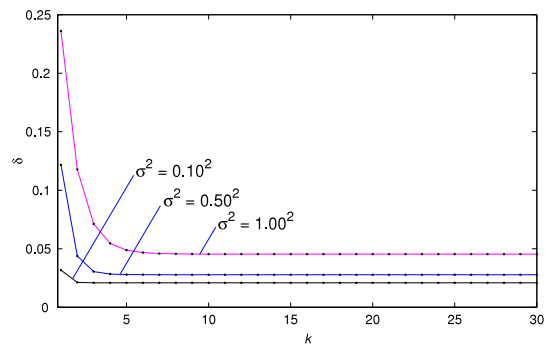
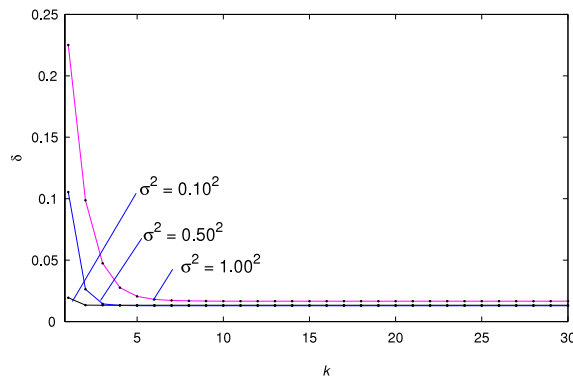
- The estimation errors  $\delta$  given by the LSI algorithm and the two-stage LSI algorithm become smaller as the data length  $L$  increases.
- Under the same data length and noise variance, the estimation accuracy of the two algorithms is close.
- As the noise-to-signal ratio becomes small, the parameter estimation errors given by the two-stage LSI algorithm become small and the parameter estimates converge fast to their true values for the same data length  $t = L$ .
- The two-stage LSI algorithm has fast convergence rates and needs only a few iterations to converge to their true values.

**Table 4**The two-stage LSI estimates and errors versus iteration  $k$  ( $L = 1000$ ).

$\sigma^2$	$k$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
$0.10^2$	1	-1.59185	0.79393	0.39630	0.31191	-0.57976	3.17396
	2	-1.60180	0.80310	0.39877	0.30737	-0.59918	2.12038
	5	-1.60215	0.80342	0.39885	0.30721	-0.59985	2.08740
	10	-1.60215	0.80342	0.39885	0.30721	-0.59985	2.08739
$0.50^2$	1	-1.49095	0.70808	0.38465	0.38643	-0.46948	12.16564
	2	-1.58094	0.78698	0.39230	0.34868	-0.57337	4.37908
	5	-1.60623	0.80916	0.39445	0.33806	-0.60257	2.78967
	10	-1.60651	0.80940	0.39447	0.33795	-0.60288	2.77809
$1.00^2$	1	-1.36445	0.60223	0.37486	0.47944	-0.34398	23.60467
	2	-1.50232	0.71847	0.38326	0.42213	-0.49121	11.77593
	5	-1.59739	0.79863	0.38905	0.38261	-0.59274	4.87865
	10	-1.60536	0.80536	0.38953	0.37929	-0.60125	4.54018
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

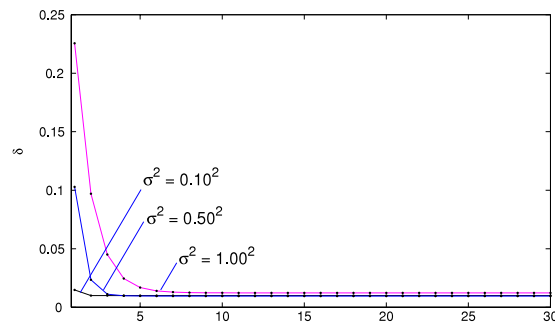
**Table 5**The two-stage LSI estimates and errors versus iteration  $k$  ( $L = 2000$ ).

$\sigma^2$	$k$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
$0.10^2$	1	-1.59450	0.79572	0.39902	0.30434	-0.60283	1.93844
	2	-1.60248	0.80284	0.39981	0.30115	-0.61406	1.33561
	5	-1.60263	0.80298	0.39983	0.30109	-0.61427	1.32626
	10	-1.60263	0.80298	0.39983	0.30109	-0.61427	1.32626
$0.50^2$	1	-1.49354	0.70788	0.39515	0.35248	-0.49772	10.53550
	2	-1.58517	0.78804	0.39836	0.31646	-0.59486	2.63224
	5	-1.60723	0.80734	0.39914	0.30780	-0.61824	1.28892
	10	-1.60739	0.80748	0.39914	0.30773	-0.61841	1.28484
$1.00^2$	1	-1.35502	0.59085	0.39167	0.41784	-0.36110	22.50624
	2	-1.50120	0.71473	0.39562	0.36030	-0.51005	9.86683
	5	-1.59664	0.79560	0.39820	0.32274	-0.60731	2.04856
	10	-1.60376	0.80163	0.39839	0.31993	-0.61456	1.66025
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

**Fig. 3.** The two-stage LSI estimation errors  $\delta$  versus  $k$  with different  $\sigma^2$  ( $L = 1000$ ).**Fig. 4.** The two-stage LSI estimation errors  $\delta$  versus  $k$  with different  $\sigma^2$  ( $L = 2000$ ).

**Table 6**The two-stage LSI estimation and errors versus iteration  $k$  ( $L = 3000$ ).

$\sigma^2$	$k$	$a_1$	$a_2$	$b_1$	$b_2$	$d_1$	$\delta$ (%)
$0.10^2$	1	-1.59450	0.79548	0.40011	0.30229	-0.61184	1.48298
	2	-1.60159	0.80194	0.40071	0.29987	-0.62055	0.99915
	5	-1.60170	0.80203	0.40072	0.29983	-0.62067	0.99376
	10	-1.60170	0.80203	0.40072	0.29983	-0.62067	0.99376
$0.50^2$	1	-1.49087	0.70455	0.40120	0.34475	-0.50657	10.28509
	2	-1.58279	0.78570	0.40309	0.30912	-0.60089	2.34527
	5	-1.60369	0.80415	0.40353	0.30103	-0.62234	0.96080
	10	-1.60382	0.80427	0.40353	0.30097	-0.62248	0.95681
$1.00^2$	1	-1.34827	0.58382	0.40465	0.40543	-0.36604	22.55197
	2	-1.49692	0.71085	0.40611	0.34616	-0.51518	9.70069
	5	-1.59223	0.79230	0.40704	0.30815	-0.61080	1.67920
	10	-1.59906	0.79814	0.40711	0.30542	-0.61766	1.22932
True values		-1.6000	0.8000	0.4000	0.3000	-0.6400	

**Fig. 5.** The two-stage LSI estimation errors  $\delta$  versus  $k$  with different  $\sigma^2$  ( $L = 3000$ ).

## 5. Conclusions

This paper presents a two-stage LSI algorithm for CARMA models. The proposed algorithm requires less computational load than the least squares based iterative algorithm. The simulation results show that the proposed algorithm has fast convergence rates and can generate highly accurate parameter estimates after only several iterations.

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